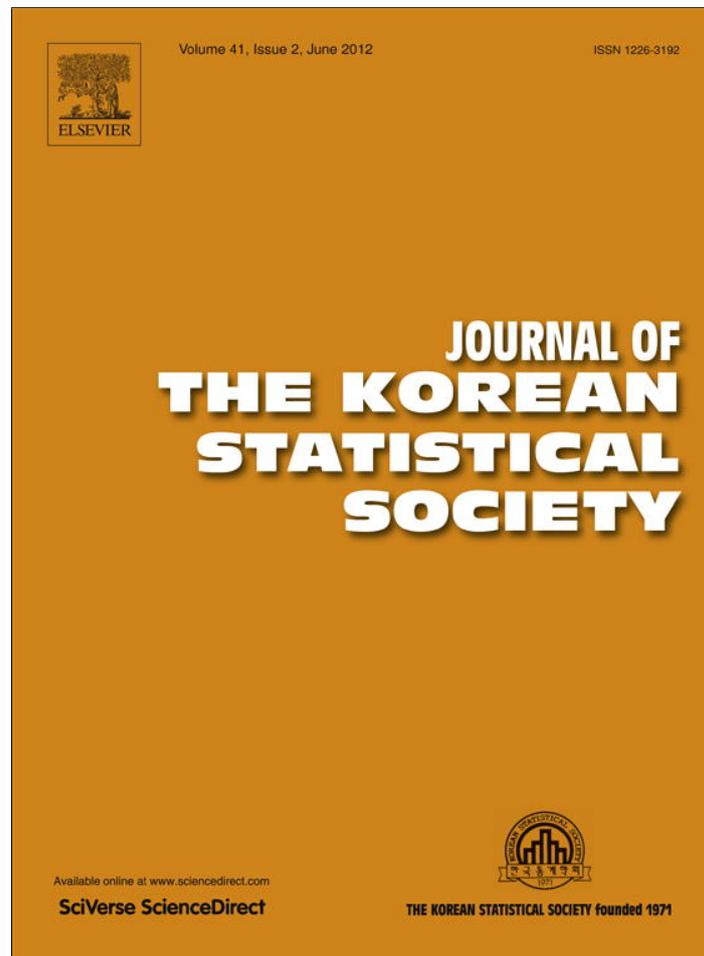


Provided for non-commercial research and education use.
Not for reproduction, distribution or commercial use.



This article appeared in a journal published by Elsevier. The attached copy is furnished to the author for internal non-commercial research and education use, including for instruction at the authors institution and sharing with colleagues.

Other uses, including reproduction and distribution, or selling or licensing copies, or posting to personal, institutional or third party websites are prohibited.

In most cases authors are permitted to post their version of the article (e.g. in Word or Tex form) to their personal website or institutional repository. Authors requiring further information regarding Elsevier's archiving and manuscript policies are encouraged to visit:

<http://www.elsevier.com/copyright>



Contents lists available at SciVerse ScienceDirect

Journal of the Korean Statistical Society

journal homepage: www.elsevier.com/locate/jkss

Discussion

Discussion: Time-threshold maps: Using information from wavelet reconstructions with all threshold values simultaneously

Hernando Ombao

Department of Statistics, University of California at Irvine, United States

ARTICLE INFO

Article history:

Received 23 February 2012

Available online 9 March 2012

I would first like to point out the broad connections of this work with those in the literature and then highlight what I believe are its major contributions.

Consider the model $Y(s_i) = f(s_i) + \epsilon(s_i)$, where f is the unknown function that has to be estimated from the data $Y(s_i)$ via a reconstruction of thresholded wavelet coefficients. One motivation for developing the time-threshold map is that using only a single threshold on a wavelet coefficient could miss potentially important features in f . A different threshold might produce a different estimate that brings out a different feature f .

This “single-threshold-per-coefficient” approach is broadly related to the standard practice in statistics, where we often either simply *assume* a model to be true or *select one* best model from a class of models using some objective criterion. We then proceed with any estimation and inference by pretending that the selected model is the correct model, that is, the selected model is the true model that generated the observed data. This approach is risky, because it ignores uncertainty in the model selection step. It also ignores the fact that the class of models is merely an approximation of the underlying truth and that each member of the class could potentially capture or explain different aspects of the underlying function.

To account for model uncertainty in statistical inference, one can apply the Bayesian model averaging (BMA) approach. Here, the final estimate is some weighted average of estimates from different models and the weights are proportional to the posterior probability of each model (see [Hoeting & Madigan, 1999](#) for a tutorial on the implementation of BMA). [Madigan and Raftery \(1994\)](#) note that BMA gives better predictive ability than using only the single selected model. Another related area is on curve aggregation (see [Bunea, Ombao, & Auguste, 2006](#)), where the final estimate is some weighted average of several curve estimates.

On a very specific problem of estimating the spectral matrix and coherence, [Fiecas and Ombao \(2011\)](#) develop an estimator which is a linear combination of some parametric estimator and a non-parametric estimator. Here, the primary estimand is a two-dimensional matrix defined on the frequency interval $(-\pi, \pi)$. The primary motivation of such an approach was that parametric estimates of the spectrum derived from an auto-regressive (AR) model with a sufficiently high order are well localized in frequency ω and hence can capture peaks that would have otherwise been flattened by a smoothing procedure. A non-parametric estimator, on the other hand, is robust to model mis-specification. Let $\hat{\mathbf{f}}_1(\omega)$ be the parametric estimator derived from some objective criterion such as the Akaike information criterion or Bayesian information criterion, and let $\hat{\mathbf{f}}_2(\omega)$ be a non-parametric estimator such as a smoothed periodogram matrix. The resulting estimator is $\hat{\mathbf{f}}(\omega) = W(\omega)\hat{\mathbf{f}}_1(\omega) + [1 - W(\omega)]\hat{\mathbf{f}}_2(\omega)$, where the frequency-specific weight for the parametric estimator is proportional to the mean-squared error (Hilbert–Schmidt norm) of the other; that is, $W(\omega) = \mathbb{E}\|\hat{\mathbf{f}}_1(\omega) - \mathbf{f}(\omega)\|_{\text{HS}}$.

One major contribution of the paper by Fryzlewicz is that it rigorously develops the analogy of “combining estimates” under the framework of wavelet estimation. In the case of wavelet estimation, a particular threshold corresponds to an optimal estimate arising from a model that assumes a particular smoothness of the underlying function f . Fryzlewicz

E-mail address: hombao@uci.edu.

introduces an interesting quantity $\Delta \mathbf{x}_\lambda = \lim_{h \rightarrow 0^+} \mathbf{x}_{\lambda+h} - \mathbf{x}_\lambda$, which indicates precise changes in the estimates as the threshold increases. This quantity could serve to identify “stable” subintervals of the domain of f from those subintervals that are sensitive to changes in the threshold.

I believe that the approach in Fryzlewicz can also be applied to obtain a “more complete” multi-resolution estimate of f . One can produce level-dependent (or frequency-band-specific) estimates of f obtained from the varying thresholds. Consequently, one can visually observe long-range patterns and highly localized patterns in f . This would be highly informative to the user who might be interested in differentiating between global patterns from transient features.

This paper opens up some new lines of research on wavelet estimation, and one area where I hope to see this work explored is in developing a principled way of combining estimates arising from different thresholds. Here, the weights may depend on the mean-squared error of the estimators and may vary across different intervals of the domain of f . For example, there may be a subinterval of the domain in which a particular set of thresholds performs better than the other sets of thresholds. I will attempt to give a more technical statement of this problem. Analogous to class of models in BMA or curve aggregation, define $\Lambda = \{\Lambda_1, \Lambda_2, \dots\}$ to be a set of all possible vectors of coefficients under consideration. Next, define $\hat{f}_\ell(s)$ to be the optimal wavelet estimator derived under the threshold set Λ_ℓ . Define the final aggregated estimate to be

$$\hat{f}(s) = \sum_{\ell} W_{\ell}(s) \hat{f}_{\ell}(s),$$

where the weights are allowed to vary across the domain s . The weights $W_{\ell}(s)$ could be inversely proportional to the mean-squared error of the corresponding estimator $\hat{f}_{\ell}(s)$. The theoretical form of the weight might be derived analytically but, in practice, it could be estimated via an application of the bootstrap procedure. This approach allows the user to identify the threshold set that gives a good (versus poor) fit at certain subintervals of the domain.

References

- Bunea, F., Ombao, H., & Auguste, A. (2006). Minimax adaptive spectral estimation from an ensemble of signals. *IEEE Transactions on Signal Processing*, 54, 2865–2874.
- Fiecas, M., & Ombao, H. (2011). The generalized shrinkage estimator for the analysis of functional connectivity of brain signals. *Annals of Applied Statistics*, 5, 1102–1125.
- Hoeting, J., Madigan, D., Raftery, A., & Volinsky, C. (1999). Bayesian model averaging: a tutorial. *Statistical Science*, 14, 382–417.
- Madigan, D., & Raftery, A. E. (1994). Model selection and accounting for model uncertainty in graphical models using Occam's window. *Journal of the American Statistical Association*, 89, 1535–1546.